

# Adolescents Who Self-Harm and Commit Violent Crime: Testing Early-Life Predictors of Dual Harm in a Longitudinal Cohort Study

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**Objective:** Self-harm is associated with violent offending. However, little is known about young people who engage in “dual-harm” behavior. The authors investigated antecedents, clinical features, and life characteristics distinguishing dual-harming adolescents from those who self-harm only.

**Methods:** Participants were from the Environmental Risk (E-Risk) Longitudinal Twin Study, a nationally representative U.K. cohort of 2,232 twins born in 1994 and 1995. Self-harm in adolescence was assessed through interviews at age 18. Violent offending was assessed using a computer questionnaire at age 18 and police records through age 22. Risk factors were assessed between ages 5 and 12. Adolescent mental health, victimization, personality functioning, and use of support services were measured at age 18.

**Results:** Self-harm was associated with violent crime (odds ratio=3.50, 95% CI=2.61–4.70), even after accounting for familial risk factors. Dual harmers had been victims of violence from childhood and exhibited lower childhood self-control

and lower childhood IQ than self-only harmers. Dual harmers experienced higher rates of concurrent psychotic symptoms and substance dependence. They also exhibited distinct personality styles characterized by resistance to change and by emotional and interpersonal lability. However, dual harmers were not more likely than self-only harmers to have contact with mental health services.

**Conclusions:** Dual harmers have self-control difficulties and are immersed in violence from a young age. A treatment—rather than punishment-oriented approach is indicated to meet these individuals’ needs. Connecting self-harming adolescents with delinquency-reduction programs and transdiagnostic approaches that target self-regulation may reduce harmful behaviors. Preventing childhood maltreatment and implementing strategies to reduce victimization exposure could mitigate risk for both internalized and externalized violence.

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Self-harm is the act of inflicting harm on oneself through the destruction of body tissue, ingestion of toxic substances, or other intentional acts, and it can occur with and without suicidal intent (1). Self-harm is a leading risk factor for suicide and a major public health problem (2–4). In the United Kingdom, where this study is based, self-harm among adolescents is of particular concern. Between 2011 and 2014, the annual incidence of self-harm increased 68% among girls ages 13–16 (3). The health care cost associated with self-harm is estimated at £162 million yearly, with the highest costs incurred by individuals under age 18 (5). High self-harm-related costs are likewise observed in the United States (6).

Some individuals who self-harm also inflict harm on others (7–11). There may be important antecedents that increase the risk of violent crime among people who self-harm. Identifying such antecedents could guide early prevention strategies

and delivery of targeted interventions to reduce interpersonal violence. However, studies have primarily examined risk factors for self-harm among violent offenders *after* they become clients of the criminal justice system (12, 13). To appropriately target assessments and treatments, clinicians need information to identify who among self-harming adolescents is at greatest risk for violent offending. Our aim in this study was to characterize the risk factors that distinguish young people who engage in both self-harm and violent crime (“dual harmers”) from those who only self-harm, using data from a nationally representative cohort of British children followed across the first two decades of life.

Our analysis capitalized on four design features. First, self-harm was assessed across adolescence and violent crime was assessed through age 22 using police records and self-reports. This allowed us to test for an association between

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self-harm and violent crime during the period when self-harm debuts (14) and criminal offending peaks (15).

Second, because the cohort comprises twins, we could conduct a co-twin-control analysis among pairs discordant for self-harm to test whether the sibling who self-harmed was more likely to offend than the one who did not self-harm. The same unmeasured risk factors that lead individuals to self-harm may also lead them to commit violent crime (16). By comparing twins who grow up in the same family, it is possible to isolate self-harm as an indicator of violent offending, independent of familial risks.

Third, the longitudinal design enabled assessment of risk factors that antedate self-harm and violent crime. Problems in self-regulation are theorized to underlie both self-harm (14, 17) and violent offending (15, 18), and they may be important targets for intervention, so we tested whether dual harmers were distinguished by low childhood self-control. In addition, we examined three risk factors identified as salient predictors of self-harm or violent crime and severe psychopathology: maltreatment, childhood self-harm behavior, and family history of psychiatric disorders (1, 15, 19). In response to external review, we also evaluated three secondary risk factors: low IQ, depression, and anxiety.

Fourth, we assessed participants' self-harm features, clinical correlates, and life characteristics. This allowed us to draw a comprehensive picture of dual harmers' psychosocial functioning. We examined self-harm method and frequency, as these are indicators of severity (4, 14, 20). We characterized dual harmers' mental health difficulties, experiences of adolescent victimization, and informant-reported personality functioning. Lastly, we evaluated their use of support services.

## METHODS

### Participants

Participants were members of the Environmental Risk (E-Risk) Longitudinal Twin Study, a birth cohort of 2,232 British children drawn from a larger register of twins born in England and Wales in 1994 and 1995 (21). Details of the E-Risk study have been reported elsewhere (22). The E-Risk sample was constructed in 1999 and 2000, when 1,116 families (93% of those eligible) with same-sex 5-year-old twins participated in home-visit assessments. In this sample, 56% of twin pairs were monozygotic and 44% were dizygotic; sex was evenly distributed within zygosity (49% male). Families were recruited to represent the U.K. population with newborns in the 1990s on the basis of residential location throughout England and Wales, as well as mother's age. Teenage mothers with twins were overselected to replace high-risk families selectively lost to the register through nonresponse. Older mothers who had twins via assisted reproduction were underselected to avoid an excess of well-educated older mothers. The study sample represented the full range of socioeconomic conditions in the United Kingdom, as reflected in families' distribution on a neighborhood-level socioeconomic index (23): 25.6% of E-Risk families live in

"wealthy achiever" neighborhoods, compared with 25.3% nationwide; 5.3% live in "urban prosperity" neighborhoods, compared with 11.6%; 29.6% live in "comfortably off" neighborhoods, compared with 26.9%; 13.4% live in "moderate means" neighborhoods, compared with 13.9%; and 26.1% live in "hard-pressed" neighborhoods, compared with 20.7%. E-Risk underrepresents "urban prosperity" households because they are likely to be childless.

Follow-up home visits took place when study participants were ages 7 (98% participation), 10 (96%), 12 (96%), and 18 (93% participation). Home visits at ages 5–12 assessed twin participants and their mothers; only twins were assessed at age 18. There were no differences between those who took part at age 18 and those who did not on socioeconomic status, as assessed when the cohort was initially defined ( $\chi^2=0.86$ ,  $p=0.65$ ), IQ score at age 5 ( $t=0.98$ ,  $p=0.33$ ), or internalizing or externalizing behavior problems at age 5 ( $t=0.40$ ,  $p=0.69$ , and  $t=0.41$ ,  $p=0.68$ , respectively).

Each twin in a twin pair was assessed by a different interviewer. Data were supplemented by searches of official records and questionnaires that are mailed, as developmentally appropriate, to teachers and co-informants nominated by participants. The Joint South London and Maudsley and Institute of Psychiatry Research Ethics Committee approved each study phase. Parents gave informed consent and twins gave assent between ages 5 and 12 and informed consent at age 18.

### Self-Harm

At age 18, participants were asked about self-harm behavior since age 12, using a life history calendar to aid recall. Ages 12–18, the years of secondary school in the United Kingdom, represent a meaningful developmental period for self-harm. Participants were asked, "Have you ever tried to hurt yourself, to cope with stress or emotional pain?" Individuals who endorsed self-harm were queried about methods. Ten behaviors were probed (e.g., cutting, burning, overdose), and participants were given the option of describing any other way they had hurt themselves. Of 2,064 participants who provided self-harm data, 280 (13.6%) reported self-harm behaviors. To assess self-harm frequency, we summed participants' responses concerning the number of times they had performed each behavior (median reported number of self-harm incidents, 6.5).

### Violent Crime

Official records of participants' criminal offending were obtained through U.K. Police National Computer record searches conducted in cooperation with the Ministry of Justice. Records include complete histories of cautions and convictions beginning at age 10, the age of criminal responsibility. Our data are complete through age 22. Violent offending was coded as a binary variable to reflect whether participants had been cautioned or convicted for a violent offense. 2,060 twins consented to record searches of their offending histories, of whom 106 (5.2%) had a record of a violent offense (see Table S1 in the online supplement).

Violent offending was also assessed via a computer questionnaire at age 18, in which participants reported on past-year offending behaviors. Violent offenses were defined to include behaviors that involved the use of force or threat of force upon a victim (e.g., robbery, assault; see Table S2 in the online supplement). Of 2,053 respondents with self-report data, 677 (33.0%) endorsed one or more violent behaviors and 338 (16.5%) endorsed two or more.

Respondents were coded as positive for violent crime if they had an official record of a violent crime or self-reported two or more violent offenses. A total of 398 of 2,051 (19.4%) participants met these criteria (see the online supplement).

### Typology of Self-Harm and Other Harm

We categorized participants into three groups for analyses: individuals coded as negative for both self-harm and violent crime (“neither harmers”;  $N=1,475$  [72.0%]), those coded as positive for self-harm and negative for violent crime (“self-only harmers”;  $N=177$  [8.6%]), and those coded as positive for both self-harm and violent crime (“dual harmers”;  $N=97$  [4.7%]).

We aimed to identify which adolescents, among those who self-harm, are most likely to commit violent crime. Therefore, self-only harmers were the comparison group of interest. However, we also conducted comparisons with adolescents who only commit violent crime (“other-only harmers”;  $N=300$  [14.6%]).

### Childhood Risk Factors

We analyzed four prespecified and theory-driven childhood risk factors: low self-control, maltreatment by an adult, childhood self-harm behavior, and family history of psychiatric disorder (1, 15, 19, 24) (see Table S3 in the online supplement). We also collected information on caregiver- and teacher-reported self-regulation difficulties at age 12 (see Table S3). In response to suggestions from external reviewers, we analyzed three secondary childhood risk factors at age 12: low IQ, depression, and anxiety (see Table S3).

### Correlates of Clinical Importance

We collected information on correlates of dual-harm behavior at age 18. We analyzed correlates in three categories with relevance for clinical practice: mental health difficulties (DSM-IV-based symptoms or diagnoses of posttraumatic stress disorder [PTSD], depression, psychosis, and substance dependence); experiences of adolescent victimization (crime victimization, maltreatment, neglect, sexual victimization, family violence, Internet/mobile-telephone victimization, and peer/sibling victimization); and informant-reported personality functioning (see Table S3).

### Service Use

At age 18, participants were queried regarding past-year treatment for emotional problems. Participants were asked whether they had used a range of services, including mental

health professionals, other supports (e.g., medical doctor, social services), and medication (see Table S3).

### Statistical Analysis

We used logistic regression to test for an association between self-harm and violent crime. We included an interaction term to test whether the association differed by sex. We used conditional logistic regression to test whether twins from discordant pairs who self-harmed were at excess risk for violent crime relative to their co-twins who did not self-harm.

We used multinomial and binomial logistic regression to predict group membership from childhood risk factors. The binomial tests were of greatest interest, as we aimed to identify the antecedents that distinguished dual from self-only harmers.

We used chi-square tests to determine whether the dual and self-only harm groups differed in the proportion of individuals reporting a high frequency of self-harm (more than 50 incidents [75th percentile of the distribution]). We used regression to test whether dual harmers were distinguished by mental health difficulties, victimization experiences, and personality functioning and to compare dual and self-only harmers on service use. Groups were included as predictors, first as a set of binary dummy codes (with the neither-harm group specified as the reference category) and then as a two-level nominal variable (to compare risk between the dual harm and self-only harm groups). We analyzed continuously distributed outcomes using ordinary least squares and binary outcomes using logistic regression.

Analyses were conducted using SAS, version 9.4 (SAS Institute, Inc., Cary, N.C.). We used survey analysis procedures to correct all analyses (except the twin-discordance analysis) for the nonindependence of twin observations by clustering standard errors at the family level. Analyses in which male and female participants were combined were adjusted for sex. Analyses were limited to individuals with complete data for self-harm and violent crime ( $N=2,049$ ); no data were imputed.

## RESULTS

Of the 2,232 participants in the original cohort, 2,066 (92.6%) were interviewed at age 18, of whom 2,049 (99.2%) had data for both self-harm and violent crime (970 of them male [47.3%]). Of the 2,049 participants included in analyses, 274 (13.4%) reported self-harm and 397 (19.4%) met criteria for violent crime.

### Association of Self-Harm and Violent Crime in Adolescence

Self-harm was more prevalent among females than males ( $\chi^2=14.93$ ,  $df=1$ ,  $p<0.001$ ), and violent crime was more prevalent among males than females ( $\chi^2=78.08$ ,  $df=1$ ,  $p<0.001$ ), but the relation between self-harm and violent offending was similar in both sexes: the odds of committing violent crime were more than three times greater for adolescents

**TABLE 1. Predicting dual versus self-only harm status from childhood risk factors in a study of adolescents who self-harm and commit violent crime<sup>a</sup>**

Childhood Risk Factor	Harm Status											
	Neither Harm (N=1,475)		Self-Only Harm (N=177)		Dual Harm (N=97)		Self-Only Versus Neither <sup>b</sup>		Dual Versus Neither <sup>b</sup>		Dual Versus Self-Only <sup>c</sup>	
	Mean	SD	Mean	SD	Mean	SD	Odds Ratio	95% CI	Odds Ratio	95% CI	Odds Ratio	95% CI
Primary risk factors												
Low self-control <sup>d</sup>	-0.19	0.94	-0.03	0.89	0.70	1.08	1.39	1.16–1.65	2.36	1.90–2.95	<b>1.82</b>	<b>1.35–2.45</b>
Family psychiatric history <sup>e</sup>	0.35	0.26	0.45	0.31	0.45	0.26	4.01	2.06–7.81	3.59	1.72–7.48	0.73	0.29–1.87
	N	%	N	%	N	%	Odds Ratio	95% CI	Odds Ratio	95% CI	Odds Ratio	95% CI
Maltreatment	58	3.9	14	7.9	18	18.6	2.28	1.09–4.76	5.33	2.91–9.76	<b>2.46</b>	<b>1.10–5.51</b>
Childhood self-harm	39	2.8	17	9.9	13	14.0	4.05	2.18–7.54	5.58	2.90–10.72	1.37	0.63–2.96
	(N=1,419)		(N=172)		(N=93)							
	Mean	SD	Mean	SD	Mean	SD	Odds Ratio	95% CI	Odds Ratio	95% CI	Odds Ratio	95% CI
Secondary risk factors												
IQ <sup>f</sup>	100.15	15.08	98.06	15.23	91.94	16.88	0.99	0.98–1.01	0.97	0.95–0.98	<b>0.98</b>	<b>0.96–0.996</b>
	N	%	N	%	N	%	Odds Ratio	95% CI	Odds Ratio	95% CI	Odds Ratio	95% CI
Depression	26	1.8	14	8.2	14	15.1	5.02	2.52–9.99	9.42	4.73–18.76	1.94	0.84–4.48
	(N=1,441)		(N=171)		(N=93)							
Anxiety	74	5.1	19	11.1	10	10.8	2.07	1.19–3.60	2.46	1.21–4.99	1.07	0.42–2.67
	(N=1,442)		(N=171)		(N=93)							

<sup>a</sup> Primary risk factors were prespecified. Secondary risk factors were added in response to peer review. Measures were assessed between ages 5 and 12. Ns for participants with data are reported when lower than the group sample size. All regression models controlled for sex. Estimates in boldface indicate a significant difference between the dual harm and self-only harm groups, which was the test of interest.

<sup>b</sup> Odds ratios are from multinomial logistic regression models.

<sup>c</sup> Odds ratios are from binomial logistic regression models.

<sup>d</sup> The self-control factor score was standardized to have a mean of 0 and a standard deviation of 1; higher scores indicate lower levels of self-control (more self-control difficulties).

<sup>e</sup> Indicates the proportion of a participant's relatives with a psychiatric disorder. Ns were 1,441, 171, and 93 participants with data on family psychiatric history in the neither-harm, self-only harm, and dual harm groups, respectively.

<sup>f</sup> Ns were 1,442, 171, and 93 participants with data on IQ in the neither-harm, self-only harm, and dual harm groups, respectively.

who self-harmed than for those who did not (males: odds ratio=3.77, 95% CI=2.46–5.78; females: odds ratio=3.27, 95% CI=2.17–4.94) (see Figure S1 in the online supplement). Therefore, male and female participants were combined in analyses (but we controlled for sex; odds ratio=3.50, 95% CI=2.61–4.70). The association remained significant when only police records for violent crime were used (odds ratio=3.26, 95% CI=2.08–5.12) and when only self-reports were used (odds ratio=3.50, 95% CI=2.57–4.76), indicating that the findings do not simply reflect common method variance.

**Effect of Familial Risk Factors on the Association Between Self-Harm and Violent Crime**

Twins who self-harmed were more likely to commit violent crime than their co-twins who did not self-harm (dizygotic twins: odds ratio=2.57, 95% CI=1.07–6.16; monozygotic twins: odds ratio=4.00, 95% CI=1.34–11.97) (see Figure S2 in the online supplement), indicating that the relation between

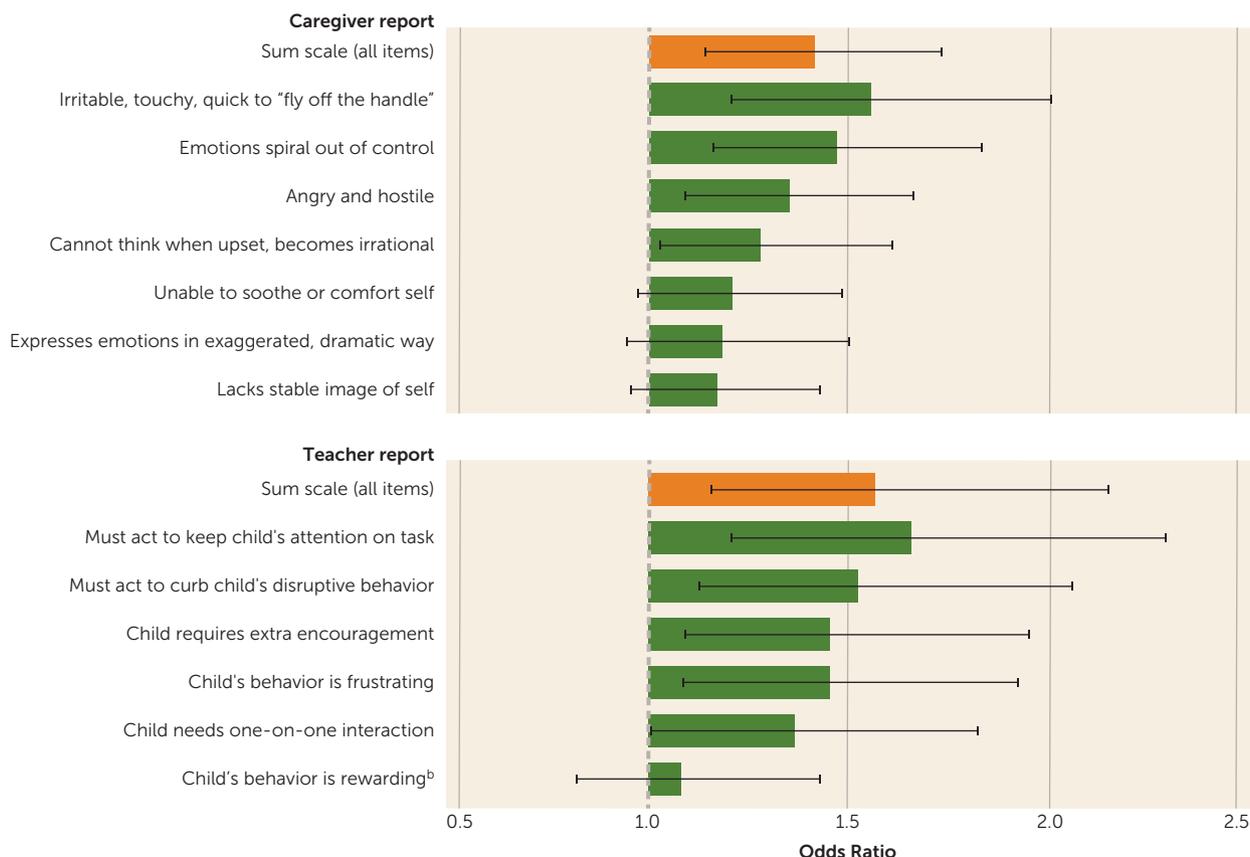
self-harm and violent offending could not be explained entirely by familial risk factors (genetics or rearing environment).

**Childhood Risk Factors Distinguishing Dual from Self-Only Harmers**

Analyses of primary risk factors showed that low childhood self-control and maltreatment predicted increased odds of being a dual versus a self-only harmer (self-control: odds ratio=1.82, 95% CI=1.35–2.45; maltreatment: odds ratio=2.46, 95% CI=1.10–5.51) (Table 1). Together, the four primary risk factors predicted membership in the dual harm relative to the self-only harm group with high accuracy (area under the curve=0.75, 95% CI=0.69–0.82, indicating a large effect [25] that requires out-of-sample replication; see the online supplement).

Analyses of secondary risk factors indicated that higher childhood IQ predicted decreased odds of being a dual versus a self-only harmer (odds ratio=0.98, 95% CI=0.96–0.996)

**FIGURE 1. Lack of self-regulation across settings in a study of adolescents who self-harm and commit violent crime<sup>a</sup>**



<sup>a</sup> The graphs indicate the odds of being in the dual harm group compared with the self-only harm group as a function of caregiver reports of children's self-regulation difficulties (top panel) and teacher reports of children's self-regulation difficulties in the classroom (bottom panel). For each item pertaining to self-regulation within the classroom, teachers were asked to rate how frequently they needed to intervene with the child. Sum scales and individual items were standardized to have a mean of 0 and a standard deviation of 1. Ns were 265 for caregiver reports and 215–216 for teacher reports. Error bars indicate 95% confidence intervals.

<sup>b</sup> Responses to this item were reverse-scored.

(Table 1). Dual harmers did not differ from self-only harmers in rates of childhood depression or anxiety.

**Dual Harmers' Self-Regulation Difficulties Across Settings**

Dual harmers' self-regulation difficulties were observable across settings. Children rated by caregivers and teachers as having more self-regulation difficulties were more likely to be in the dual harm than the self-only harm group as adolescents (caregivers' scale score: odds ratio=1.41, 95% CI=1.14–1.74; teachers' scale score: odds ratio=1.56, 95% CI=1.15–2.13) (Figure 1).

**Clinical Features and Life Characteristics Distinguishing Dual From Self-Only Harmers**

Dual and self-only harmers reported similar rates of high-frequency self-harm (more than 50 incidents; dual harm, 26.6%; self-only harm, 24.6%;  $\chi^2=0.13$ ,  $df=1$ ,  $p=0.72$ ). Given the small sample sizes for some self-harm methods, we did not conduct tests of group differences for each method. However, inspection of Figure 2A suggests that dual harmers

exhibited higher-lethality behaviors (hanging, drowning) and aggressive acts (hitting oneself or an object, banging one's head against a wall), while self-only harmers tended to engage in lower-lethality methods (cutting, scratching).

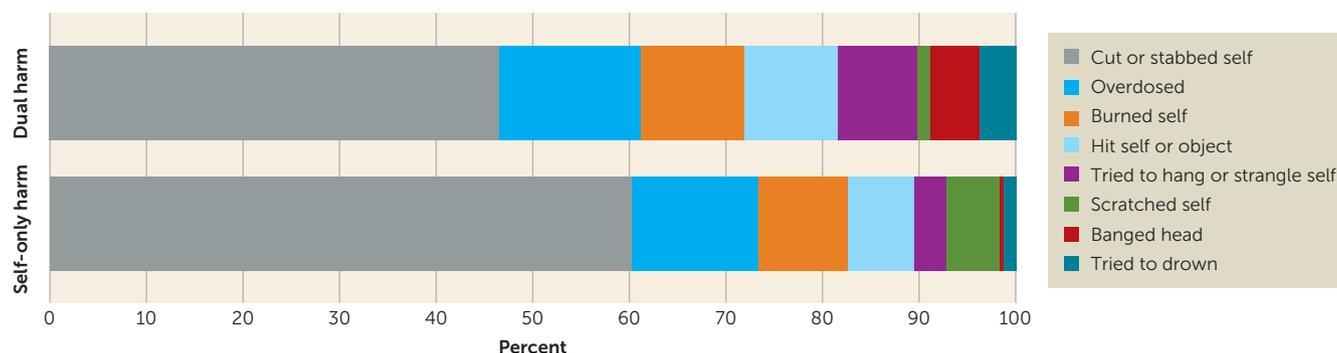
Dual harmers did not differ from self-only harmers in their risk of developing PTSD or depression. However, they were distinguished by a higher prevalence of psychotic symptoms (odds ratio=2.35, 95% CI=1.11–4.95). They were also more likely to meet criteria for alcohol dependence (odds ratio=3.29, 95% CI=1.65–6.57) and cannabis dependence (odds ratio=4.31, 95% CI=1.91–9.76) (Table 2).

Dual harmers were more likely than self-only harmers to have experienced multiple types of victimization during adolescence (polyvictimization; odds ratio=2.40, 95% CI=1.30–4.42) (see Figure S3 in the online supplement) as well as crime, maltreatment, neglect, and family violence (Table 2).

Dual harmers' personality styles were different from those of self-only harmers. Dual harmers were distinguished by greater resistance to change (lower openness; Cohen's  $d=-0.41$ ), poorer impulse control (lower conscientiousness;

**FIGURE 2. Comparing dual harmers and self-only harmers on clinical features in a study of adolescents who self-harm and commit violent crime<sup>a</sup>**

**A. Self-Reported Self-Harm Behaviors**



**B. Personality Inventory**



<sup>a</sup> Panel A displays the proportion of 393 total reported self-harm behaviors attributable to different self-harm methods. The totals were derived by summing the number of individuals who endorsed each self-harm method. Dual harmers endorsed 157 behaviors, and self-only harmers endorsed 236 behaviors. Participants were allowed to endorse multiple behaviors and could be included more than once within the total. Only behaviors endorsed by at least 3% of the sample are depicted. Panel B displays Big Five Inventory profiles, provided by one to two informants who know the participants well, for each group. Of the 1,749 participants, personality data were available for 1,730 (98.9%).

$d=-0.63$ ), and more aggressive/rude behavior (lower agreeableness;  $d=-0.46$ ). They were more outgoing (higher in extraversion); however, the effect size was modest ( $d=0.15$ ) (Figure 2B; see also Table S4 in the online supplement). Both dual and self-only harmers were more easily distressed (higher in neuroticism) and were lower in conscientiousness and agreeableness than neither harmers.

Despite their elevated rates of psychiatric comorbidity and difficult life experiences, dual harmers were not more likely than self-only harmers to be in contact with mental health professionals (psychiatrists, psychologists, counselors, or psychotherapists) or other support services (Figure 3).

**Comparisons With Other-Only Harmers**

Compared with participants who committed violent crime only, dual harmers exhibited higher rates of childhood self-harm and childhood depression, had higher rates of all adolescent mental health difficulties, were more likely to have experienced polyvictimization and nearly all types of victimization, and were lower in conscientiousness and higher in neuroticism (see Tables S5–S7 in the online supplement).

**DISCUSSION**

This study shows that self-harm and violent crime co-occur in a longitudinal population-representative contemporary cohort of British twins. The association is evident in police records and self-reports of offending. This finding is consistent with research employing population-based samples from other countries (7–10).

This study advances knowledge in five ways. First, using a co-twin-control design, we showed that the relation between self-harm and violent crime is not solely attributable to shared genetic risk or family background; self-harm itself may be an indicator of violence against others.

Second, we demonstrated that dual harmers are distinguished from self-only harmers by poor childhood self-control, including deficits in executive functioning, as indicated by lower childhood IQ. Prospective assessment enabled measurement of self-control, cognitive ability, and other antecedents prior to the onset of self-harm and criminal offending and ensured that there were no ascertainment or recall biases. Moreover, dual harmers' self-regulation difficulties were reported by multiple informants, suggesting

**TABLE 2. Comparing dual and self-only harm groups on correlates of clinical importance in a study of adolescents who self-harm and commit violent crime<sup>a</sup>**

Measure	Harm Status						Self-Only Versus Neither		Dual Versus Neither		Dual Versus Self-Only	
	Neither Harm (N=1,475) <sup>b</sup>		Self-Only Harm (N=177) <sup>c</sup>		Dual Harm (N=97) <sup>d</sup>		Odds Ratio	95% CI	Odds Ratio	95% CI	Odds Ratio	95% CI
	N	%	N	%	N	%						
<b>Mental health difficulties</b>												
PTSD	26	1.8	24	13.6	13	13.5	7.67	4.10–14.34	9.90	4.86–20.14	1.07	0.51–2.21
Depression	182	12.4	95	54.0	58	59.8	7.65	5.34–10.95	11.76	7.58–18.26	1.35	0.79–2.31
Psychotic symptoms	18	1.2	16	9.0	16	16.5	7.63	3.88–15.01	16.63	7.98–34.64	<b>2.35</b>	<b>1.11–4.95</b>
Alcohol dependence	144	9.8	23	13.0	33	34.4	1.47	0.88–2.44	4.69	2.93–7.52	<b>3.29</b>	<b>1.65–6.57</b>
Cannabis dependence	19	1.3	11	6.2	25	25.8	5.93	2.55–13.76	25.11	12.56–50.19	<b>4.31</b>	<b>1.91–9.76</b>
<b>Victimization experiences<sup>e</sup></b>												
Polyvictimization <sup>f</sup>	36	2.4	38	21.5	32	33.3	9.91	6.05–16.25	22.09	12.51–39.03	<b>2.40</b>	<b>1.30–4.42</b>
Conventional crime	179	12.1	59	33.3	54	55.7	4.08	2.82–5.91	8.77	5.72–13.44	<b>2.23</b>	<b>1.32–3.79</b>
Maltreatment	15	1.0	17	9.6	20	20.6	10.13	4.79–21.45	25.63	11.86–55.43	<b>2.63</b>	<b>1.19–5.82</b>
Neglect	12	0.8	13	7.3	14	14.6	8.95	4.05–19.80	22.05	9.97–48.75	<b>2.30</b>	<b>1.01–5.25</b>
Sexual	7	0.5	25	14.1	15	15.6	27.94	11.62–67.18	52.17	20.52–132.68	1.79	0.83–3.86
Family	118	8.0	41	23.2	33	34.4	3.39	2.24–5.13	6.12	3.79–9.89	<b>1.99</b>	<b>1.11–3.57</b>
Internet	69	4.7	24	13.6	15	15.6	2.60	1.57–4.31	4.72	2.46–9.09	1.99	0.88–4.50
Peer	152	10.3	64	36.2	36	37.5	4.54	3.17–6.50	5.65	3.54–9.02	1.19	0.68–2.09

<sup>a</sup> Measures were assessed at age 18. All regression models controlled for sex. Estimates in boldface indicate a significant difference between the dual harm and self-only harm groups, which was the test of interest. Prevalence estimates are derived using the number of participants with data for the measure; this was occasionally slightly lower than the group sample size.

<sup>b</sup> Number of participants with data ranged from 1,473 to 1,475.

<sup>c</sup> Number of participants with data ranged from 176 to 177.

<sup>d</sup> Number of participants with data ranged from 96 to 97.

<sup>e</sup> Prevalences for victimization experiences indicate the percentage of individuals who reported a severe level of exposure.

<sup>f</sup> Polyvictimization is defined here as three or more types of victimization.

that this early-emerging risk factor is observable across settings. In addition to their experiences of childhood dysregulation, dual harmers were characterized in adolescence by a triad of personality features that typifies emotional and interpersonal lability: low conscientiousness, low agreeableness, and high neuroticism (26). (Dual and self-only harmers did not differ on neuroticism, but this trait was elevated in both groups.) Apparently, dual harmers' self-control difficulties are a stable core feature of their personalities.

Third, we showed that dual harmers are differentiated from self-only harmers by a history of childhood maltreatment. Furthermore, dual harmers were more likely to have been exposed to adolescent victimization. More than 80% of dual harmers had experienced at least one type of victimization, and one-third had experienced polyvictimization. These findings signal a need for primary and secondary preventive strategies to reduce continuity in victimization among individuals at risk for dual-harm behavior.

Fourth, we found that dual harmers are distinguished from self-only harmers by higher rates of psychotic symptoms, alcohol dependence, and cannabis dependence. A previous analysis (7) did not find differences in risk for cannabis-related problems between dual and self-only harmers. However, that study employed a retrospective survey of adults and DSM-5-based lifetime diagnoses. Ours is the first study, to our knowledge, to test these associations within a prospective sample and to demonstrate the role of psychosis in the self/other harm typology. Dual harmers suffer

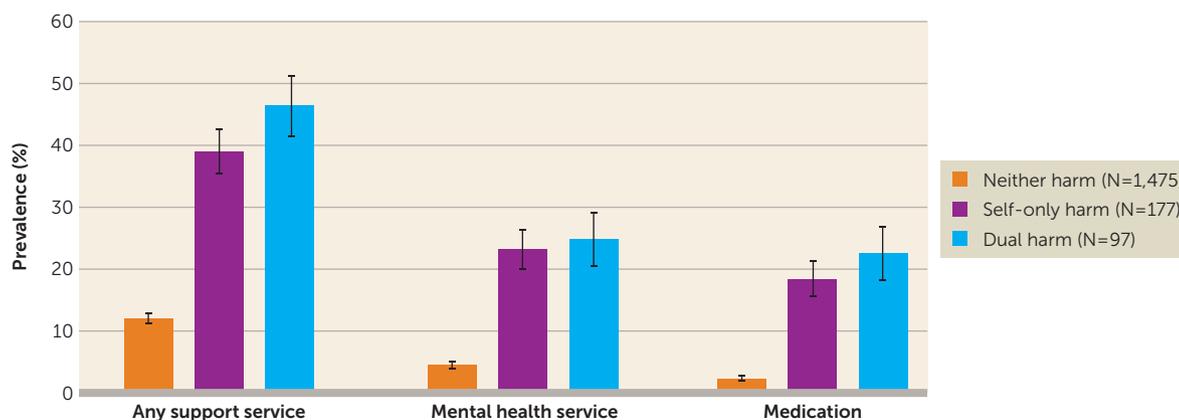
significant psychiatric comorbidity; comprehensive diagnostic assessment is needed to appropriately target interventions within this population.

Lastly, we found that dual harmers were not more likely than self-only harmers to encounter mental health services. Recent U.K.-based data (3, 27) suggest long waiting lists and high thresholds in accessing treatment, and similar challenges exist in the United States (28). Research on hospital- and community-based youth violence prevention services identifies mistrust of authorities as a barrier to treatment engagement (29). Dual harmers' psychosocial difficulties and prior experiences with the juvenile justice system may impede service use.

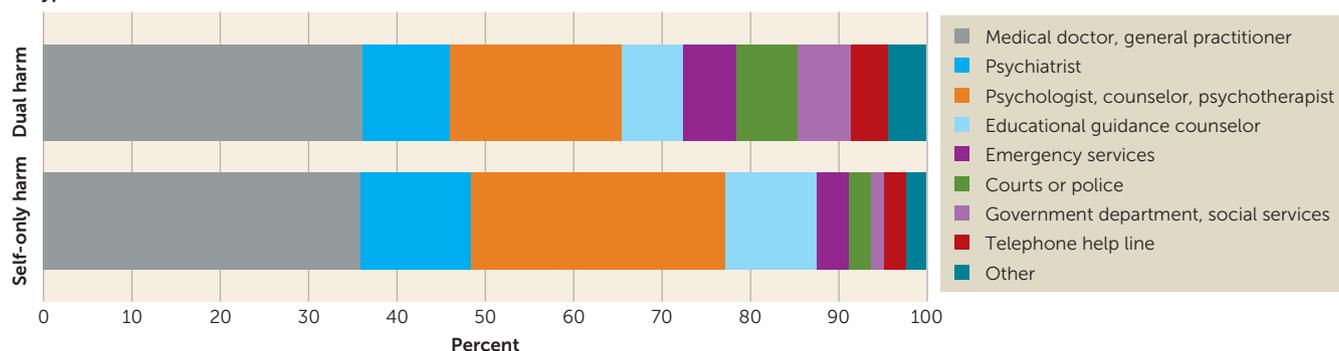
This study has limitations. First, the sample comprised twins, and the results may not generalize to singletons. However, the prevalences of antisocial behavior and mental health problems are similar for twins and singletons (30, 31), and the association between self-harm and violent crime has been documented in non-twin samples (7–10). Second, participants were followed only to the beginning of young adulthood. Future research will determine whether the findings pertain to older age groups. Third, results may vary with historical and cross-national differences in crime-control policy. Fourth, findings concerning risk factors require replication. However, primary antecedents were selected on the basis of prior theoretical and empirical evidence, increasing the likelihood of replication. Fifth, our co-twin-control analyses included a rather small number of

**FIGURE 3. Comparing dual harmers and self-only harmers on past-year service use at age 18 in a study of adolescents who self-harm and commit violent crime<sup>a</sup>**

**A. Service and Medication Use**



**B. Types of Services Used**



<sup>a</sup> Panel A displays the prevalence of any service use, mental health service use, and medication use across the groups. There were no significant differences between the self-only and dual harm groups (all p values >0.10). Error bars represent robust standard errors. Panel B displays the proportion of 236 total reported services used attributable to different service types; totals were derived by summing the number of individuals who reported using each service type in the past year. Dual harmers reported using 101 services, and self-only harmers reported using 135 services. Participants were allowed to report multiple services and could be included more than once within the total. Only services reported by at least 3% of the sample are depicted.

“informative cases” (pairs discordant for violent offending). The results will need to be replicated in samples with a higher prevalence of discordant pairs. Sixth, we designed our assessments of self-harm and violent offending consistent with recommendations for best practice. However, differences in the types of assessment methods used across constructs may have affected our results. Finally, we are limited in our ability to infer causality. Assessment of self-harm and crime spanned much of the same period. Furthermore, within-twin-pair associations between self-harm and violent offending may be confounded by twin-specific environmental differences. Additionally, our research can only support low childhood self-control, low IQ, and maltreatment as indicators of risk for dual-harm behavior, not necessarily indicators of causation. Establishing whether associations are causal, however, is secondary to this study’s primary aim of informing mental health treatment.

This study has a number of implications. First, given the robust link between self-harm and harm toward others, research on self-harm—even when conducted in community

samples, not only in clinical or forensic settings—should collect data on interpersonal violence. Second, theoretical models of self-harm can generate testable hypotheses for research on dual-harm behavior. Many theories propose that self-harm serves an emotion-regulatory function (14, 17). Recently developed models hold that several proximal risk factors lower perceived barriers to initiating self-harm; however, the affective benefits of self-harm are its primary maintaining factor (32, 33). These benefits may lower barriers to engagement in other harmful behaviors, including violent crime. Although our study did not directly test this question, our findings support further investigation of self-regulation as a mediating factor. The interpersonal theory of suicide (34) posits that self-harm increases risk for suicide by habituating individuals to the fear and pain associated with harming oneself. Such habituation may also increase risk for harming others, or it may occur through repeated aggression toward others. Research on the mechanisms underlying dual harm presents opportunities for interdisciplinary collaboration. Self-harm and offending have largely been studied separately within the fields

of psychology, psychiatry, and criminology; collaborative cross-talk can inform more effective preventions and treatments.

Third, clinical guidelines recommend evaluation of risk for suicide following self-harm (35, 36). Our results support a recommendation of assessment of risk for violence toward others as well, particularly when the clinical picture comprises relevant antecedents and correlates. Furthermore, dual-harming prisoners should be closely monitored for suicidal behavior. Fourth, improving self-control among self-harmers could help prevent violent crime. Self-control training has been shown to reduce delinquency (37) and could be delivered to patients who self-harm. In addition, dual harmers often experience psychiatric comorbidity. Transdiagnostic approaches that target self-regulation (e.g., mindfulness-based approaches for emotion regulation) may reduce harmful behaviors and co-occurring psychopathology (38). Lastly, our findings support recommending the application of available interventions to prevent childhood maltreatment (39) as well as implementation of exposure-reduction strategies (e.g., education on self-protective measures) and evidence-based programs (40, 41) to prevent revictimization in adolescence. Dual harmers have been immersed in violence from a young age; a treatment-oriented rather than a punishment-oriented approach is indicated to meet these individuals' needs. Such an approach could also yield substantial reductions in violent offending: one in four other-harming adolescents was a dual harmer in this population-representative study.

There is a pressing demand for improvements in adolescent mental health services (42) and psychological treatment research (43, 44). Our analysis responds to this demand by identifying several opportunities for early-years prevention and intervention science (43, 44). Connecting vulnerable adolescents with delinquency-reduction programs that target self-control, prevention of maltreatment and victimization, and improvement in children's self-regulation abilities could significantly reduce the health and social burdens attributable to internalized and externalized violence.

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